

Short Communication

## SOME REMARKS ON THE STABILITY OF THE CHARACTERIZATION OF THE COMPOSED RANDOM VARIABLES

NGUYEN HUU BAO

Let  $\xi$  be a random variable (r. v.) with the characteristic function  $\varphi(t)$  and  $\eta$  be a r. v. with the generating function  $a(z)$ . It is known [2] that the composed r. v. of the two random variables  $\xi, \eta$  denoted by  $\langle \xi, \eta \rangle$  is a r. v. with the characteristic function  $\psi(t) = a(\varphi(t))$ . In [1] we have dealt with the class  $N$  consisting of the characteristic functions of composed r. v.'s and proved that it is a proper extension of the class  $L$  of characteristic function of infinitely divisible laws. However, there are still many properties of the class  $N$  which have not yet considered.

In this paper we shall investigate a subclass  $N_\epsilon$  of  $N$  possessing the stability in the following sense: the small changes in the distribution functions of the components  $\xi$  and  $\eta$  only lead to the small changes in the distribution function of the composed r. v.  $\langle \xi, \eta \rangle$ . We will give some remarks on stability conditions for this class.

1. Suppose that  $\psi_1(t)$  and  $\psi_2(t)$  are two characteristic functions of the class  $N_\epsilon$  with the same generating function:

$$\psi_1(t) = a(\varphi_1(t)), \quad \psi_2(t) = a(\varphi_2(t)), \quad (1)$$

where  $a(z)$  satisfies the following conditions:

$$|a(z_1) - a(z_2)| \leq K |z_2 - z_1| \quad (2)$$

for all complex numbers  $z_1, z_2, |z_1| \leq 1, |z_2| \leq 1$ , and  $K$  is a constant.

If for any sufficiently small  $\epsilon$  ( $0 < \epsilon < 1$ ), we can choose a number  $T = T(\epsilon)$  ( $T(\epsilon) \rightarrow +\infty$  when  $\epsilon \rightarrow 0$ ) so that

$$\max_{|t| \leq T(\epsilon)} |\varphi_1(t) - \varphi_2(t)| \leq \epsilon \quad (3)$$

then we shall have the following estimation

$$\lambda(\Psi_1; \Psi_2) \leq \max \left\{ K\varepsilon; \frac{1}{T(\varepsilon)} \right\}, \quad (4)$$

where  $\Psi_1(x)$  and  $\Psi_2(x)$  are two distribution functions which has the corresponding characteristic functions  $\psi_1(t)$ ,  $\psi_2(t)$  and the metric  $\lambda$  is defined by

$$\lambda(\Psi_1; \Psi_2) = \min_{T>0} \max \left\{ \max_{|t| \leq T} |\psi_1(t) - \psi_2(t)|; \frac{1}{T} \right\} \quad (5)$$

Indeed, from (1), (2) and (3), we have, for  $|t| \leq T(\varepsilon)$ ,

$$|\psi_1(t) - \psi_2(t)| = |a(\varphi_1(t)) - a(\varphi_2(t))| \leq K\varepsilon,$$

and hence (4) follows, by the definition of  $\lambda$ .

2. If  $\nu$  is the r. v. having the Poisson law with the parameter  $\lambda > 0$  and  $\varphi_1(t)$  is the characteristic function of the r. v.  $\xi$  having  $\varepsilon$ -exponential law (i. e.  $\exists T(\varepsilon)$ ,  $T(\varepsilon) \rightarrow +\infty$  when  $\varepsilon \rightarrow 0$ ,  $\forall t$ ,  $|t| \leq T$ ,  $|\varphi_1(t) - \frac{1}{1-it\varepsilon}| \leq \varepsilon$ ) then the composed r. v. of  $\xi$  and  $\nu$  has the distribution function  $\Psi_1(x)$  which satisfies the following estimation

$$\lambda(\Psi_1; \Psi_1^{\lambda\theta}) \leq \max \left\{ e^{4\lambda} \varepsilon; \frac{1}{T(\varepsilon)} \right\}, \quad (6)$$

where  $\Psi_1^{\lambda\theta}(x)$  is the distribution function with the characteristic function

$$\psi_1^{\lambda\theta}(t) = e^{\lambda(\frac{1}{1-it\varepsilon} - 1)}.$$

3. If  $\nu$  is r. v. having the Poisson law with the parameter  $\lambda > 0$  and  $\xi$  has the  $\varepsilon$ -Normal distribution function (i.e. its the characteristic function  $\varphi_2(t)$  satisfies the estimation:  $|\varphi_2(t) - e^{-t^2/2}| \leq \varepsilon$ ,  $\forall t: |t| \leq T(\varepsilon)$ ,  $T(\varepsilon) \rightarrow +\infty$  when  $\varepsilon \rightarrow 0$ ). Then the composed r. v. of  $\nu$  and  $\xi$  has the distribution function  $\Psi_2(x)$  which satisfies the following estimation:

$$\lambda(\Psi_2; \Psi_2^{0,1,\lambda}) \leq \max \left\{ e^{4\lambda} \varepsilon; \frac{1}{T(\varepsilon)} \right\}, \quad (7)$$

where  $\Psi_2^{0,1,\lambda}(x)$  is the distribution with the characteristic function

$$\psi_2^{0,1,\lambda}(t) = e^{\lambda(e^{-t^2/2} - 1)}.$$

4. If  $\nu$  is r.v. having the binomial distribution function with the parameters  $p$ ,  $q$  and  $\xi$  has the  $\varepsilon$ -exponential distribution, then the composed r.v. of  $\nu$  and  $\xi$  has the distribution function  $\Psi_3(x)$  which satisfies the following estimation:

$$\lambda(\Psi_3; \Psi_3^{p,\theta}) \leq \max \left\{ np(1+2p)^{n-1}\varepsilon; \frac{1}{T(\varepsilon)} \right\}, \quad (8)$$

where  $\Psi_3^{p,\theta}$  is the distribution with the characteristic function

$$\psi_3^{p,\theta}(t) = \left[ 1 + p \left( \frac{1}{1 - i\theta t} - 1 \right) \right]^n.$$

5. If  $\nu$  is the r.v. having the negative binomial distribution with the parameters  $p$ ,  $q$  and  $\xi$  has the characteristic  $\varphi_4(t)$  which is  $\varepsilon$ -exponential then the composed r.v. of  $\nu$  and  $\xi$  has the distribution function  $\Psi_4(x)$  satisfying the following estimation

$$\lambda(\Psi_4; \Psi_4^{p,q,\theta}) \leq \max \left\{ \frac{p}{q}\varepsilon; \frac{1}{T(\varepsilon)} \right\}, \quad (9)$$

where  $\Psi_4^{p,q,\theta}(x)$  is the distribution function with the characteristic function:

$$\psi_4^{p,q,\theta}(t) = \frac{p - i\alpha t}{p - i\theta t} \quad (\alpha = \theta q).$$

The above Remarks 2, 3, 4, 5 are immediate from Remark 1 since the corresponding generating functions clearly satisfy the condition (2). Indeed, for instance, to show Remark 5, let  $a_4(z)$  be the generating function of the negative-binomial law, i.e.:

$$a_4(z) = p[1 - qz]^{-1}.$$

For any the complex numbers  $z_1$ ,  $z_2$  satisfying  $|z_1| \leq 1$ ,  $|z_2| \leq 1$ , we have the following estimation:

$$\begin{aligned} |a_4(z_1) - a_4(z_2)| &= \left| \frac{p}{1 - qz_1} - \frac{p}{1 - qz_2} \right| \\ &\leq \frac{pq|z_1 - z_2|}{|1 - qz_1| \cdot |1 - qz_2|}. \end{aligned} \quad (10)$$

Notice that

$$\begin{aligned} |1 - qz_1| &\geq |1 - q|z_1|| \geq 1 - q \text{ for all } |z_1| \leq 1, \\ |1 - qz_2| &\geq |1 - q|z_2|| \geq 1 - q \text{ for all } |z_2| \leq 1. \end{aligned} \quad (11)$$

From (10) and (11) it follows that

$$|a_4(z_1) - a_4(z_2)| \leq \frac{pq|z_1 - z_2|}{(1 - q)^2}.$$

Thus  $a_4(z)$  satisfies the condition (2) with the constant  $K = \frac{p}{q}$ .

## REFERENCES

1. Nguyen Huu Bao, *On the stability of the characterization of the distribution function*, doctor thesis, Hanoi 9 - 1989.
2. William Feller, *An introduction to probability theory and its applications*, New York - London - Sydney, 1966.

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Hydrological Faculty,  
Hanoi Water Resources College,  
Hanoi, Vietnam.